# CYP-C Data Analysis Using SAS II 

CYP-C Research Champion Webinar<br>November 24, 2017<br>Jason D. Pole, PhD



PEDIATRIC ONCOLOGY GROUP OF ONTARIO

## Overview

- Data Analysis
- Introduction to time-to-event analysis
- Kaplan-Meier Curves
- Cumulative Incidence Curves
- Introduction to Cox Proportional Hazards Modeling


## Introduction to Time-To-Event Analysis

## Time-To-Event (TTE) Analysis

- Often referred to as survival analysis
- Modelling technique allows you to examine the occurrence and timing of any event
- Time has two components
- Scale
- Years, days, hours, minutes, seconds, microseconds
- Selection of scale has little impact on analysis
- Only effects the intercept
- Origin
- Often implicit but can have large effect on estimates
- We use time of diagnosis as Time $=0$ but really we really want time of disease onset
- Time of diagnosis is affected by so many things
- Age, sex, access to care, symptoms etc.
- In RCTs it is time of randomization


## Time-To-Event (TTE) Analysis II

- Interested in the frequency of events happening over a period of observation
- By counting frequency over time we can think of this as the density of events



## Censoring

- Describes periods of no observation
- Many different kinds of censoring
- Left - some period before you start observing where events could occur
- Right - some period after you stopped observing where events could occur
- Interval - combines both left and right censoring



## Describing TTE Distributions

- Cumulative Distribution Function
- Tells us the probability that the variable T will be less than or equal to any value of time ( $t$ ) we choose $\mathrm{F}(\mathrm{t})$
- Survival Function
- Probability of surviving beyond $t$
$-S(t)=1-F(t)$
- $S(t)$ is a probability: bounded by 0 and 1


## Describing TTE Distributions 2

- Hazard Function
- Quantifies the instantaneous risk that an event will occur at time $t$
- We condition this on having survived to time $t$
- Describes the number of events per interval of time
- The survival function and hazard function are all equivalent ways describing a continuous probability function


## Data Structure in TTE

- For basic TTE analysis (no left censoring)
- For each unit of analysis
- time from start of observation (origin) to event or censoring (measure in any scale you choose)
- Status at end of time (often called censor)
- Status = 0 = person had event (observed event)
- Status = 1 = person was censored (observation ended)


| Subject | Time | Censor | Tx_Arm | Age |
| :---: | :---: | :---: | :---: | :---: |
| A | 2.00 | 0 | 2 | 1 |
| B | 3.00 | 1 | 2 | 2 |
| C | 0.50 | 0 | 1 | 2 |
| D | 2.75 | 0 | 1 | 3 |
| E | 2.25 | 1 | 2 | 1 |

## Kaplan-Meier Estimator

## Kaplan-Meier Estimator

- Most widely used method to estimate the survival function
- Also known as the product-limit estimator
- In 1958, Kaplan and Meier demonstrated that this method was the nonparametric maximum likelihood estimator (although the method had been used for years earlier)


## Overall Survival

DATA T7; SET T6;
IF DUMALL = 1;
TimeLastFU = LAST_CONTACT_DATE - DX_DATE;
LABEL TIMELASTFU = 'NO. OF DAYS BETWEEN DIAGNOSIS AND LAST FU';
TimeDeath = DateDeath - DX_DATE;
LABEL TIMEDEATH = 'NO. OF DAYS BETWEEN DIAGNOSIS AND DEATH';

```
/* SETS ALL POST-MORTEM DEATHS TO DAY ZERO */
```

/* CensOS = 1 = PATIENT IS ALIVE */

If TimeDeath < 0 then TimeDeath $=0$;
If DateDeath $=$. then TimeDeath $=$.;
If TimeDeath $=$. then CensOS = 1; else CensOS = 0;
TimeSurvival = Min (TimeLastFU, TimeDeath);
RUN;
PROC LIFETEST DATA = T7;
TIME TIMESURVIVAL*CENSOS(1);
RUN;

## The LIFETEST Procedure

## Product-Limit Survival Estimates

| Time |  |  | Survival |  |  |
| ---: | ---: | :---: | :---: | :---: | :---: | :---: |
| Standard | Number | Number |  |  |  |
| Survival | Survival | Failure | Error | Failed | Left |
| 1825.00 | 0.9127 | 0.0873 | 0.00600 | 209 | 927 |

## PROC LIFETEST DATA = T7 PLOT = (S); TIME TIMESURVIVAL*CENSOS(1);

## RUN;

Product-Limit Survival Estimate


DATA T7; SET T7;
IF 0 <= DX_AGE <= 0 THEN EARLY_AGE = 'INFANT';
IF $1<=$ nX $\triangle G F<=5$ THFN FARIY $\Delta G F=$ 'YOIING'.
IF 6 <
RUN:
The LIFETEST Procedure

Summary of the Number of Censored and Uncensored Values

| Stratum | EARLY AGE | Total | Failed | Censored | Percent Censored |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | INFANT | 77 | 35 | 42 | 54.55 |
| 2 | OLD | 939 | 96 | 843 | 89.78 |
| 3 | YOUNG | 1707 | 85 | 1622 | 95.02 |
| Total |  | 2723 | 216 | 2507 | 92.07 |

NOTE: 9 observations with invalid time, censoring, or strata values were deleted.

| 3.00 | $0.99 / 1$ | 0.00234 | $0.0011 /$ | 4 | $1 / 03$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 5.00 | 0.9971 | 0.00293 | 0.00131 | 5 | 1702 |

## Product-Limit Survival Estimates



EARLY_AGE —— INFANT --- OLD -- - YOUNG

| Test of Equality over Strata |  |  |  |
| :--- | :--- | :--- | :---: |
| Test | Chi-Square | DF | Pr $>$ <br> Chi-Square |
|  |  |  |  |
| Log-Rank | 242.9273 | 2 | $<.0001$ |
| Wilcoxon | 263.9643 | 2 | $<.0001$ |
| -2Log(LR) | 124.7142 | 2 | $<.0001$ |

- Each test has different properties
- Wilcoxon is more sensitive to early times (is a weighted sum of deviations and by definition there are more observations in the early period)


## Event-Free Survival

```
DATA T7; SET T6;
IF DUMALL = 1;
TimeRelapse = RX_DATE1 - DX_DATE;
LABEL TIMERELAPSE = 'NO. OF DAYS BETWEEN DIAGNOSIS AND FIRST RELAPSE';
TimeLastFU = LAST_CONTACT_DATE - DX_DATE;
LABEL TIMELASTFU = 'NO. OF DAYS BETWEEN DIAGNOSIS AND LAST FU';
TimeDeath = DateDeath - DX_DATE;
LABEL TIMEDEATH = 'NO. OF DAYS BETWEEN DIAGNOSIS AND DEATH';
/* SETS ALL POST-MORTEM DEATHS TO DAY ZERO */
IF TimeDeath < 0 then TimeDeath = 0; if DateDeath = . then TimeDeath = .;
/* DEFINES EFS USING RELASPE, DEATH AND LAST FOLLOW-UP */
TimeEvent = Min(TimeRelapse, TimeLastFU, TimeDeath);
LABEL TIMEEVENT = 'NO. OF DAYS BETWEEN DIAGNOSIS AND EFS EVENT';
/* IF PATIENT DID NOT RELASPE AND DID NOT DIE THEN CENSORED */
IF (TIMERELAPSE = . AND TIMEDEATH = .) THEN CensEFS = 1; else CensEFS = 0;
IF 0 <= DX_AGE <= 0 THEN EARLY_AGE = 'INFANT';
IF 1 <= DX_AGE <=5 THEN EARLY_AGE = 'YOUNG';
IF 6 <= DX_AGE THEN EARLY_AGE = 'OLD';
RUN;
```


## PROC LIFETEST DATA = T7 PLOT = (S); <br> TIME TIMEEVENT*CENSEFS(1); <br> STRATA EARLY_AGE;

## RUN;

## Product-Limit Survival Estimates


EARLY_AGE - INFANT - - OLD $-\cdots-$ YOUNG

## Cumulative Incidence

## Cumulative Incidence

- probability that a particular event, such as occurrence of a particular disease, has occurred before a given time
- In situation with only right censoring equivalent to 1-survival
- In SAS 9.4 can be estimated using PHREG procedure, prior need to use macro


## DATA T8; SET T7;

IF TIMESURVIVAL = 0 THEN TIMESURVIVAL = 0.005; RUN;

```
%CumIncid(data=t8,
        out=CumInc,
        time=timesurvival,
        status=censos,
        event=0,
        compete=2,
        censored=1,
        strata=,
        alpha=.05,
        options=noprint plotcI);
```

RUN;

Cumulative Incidence Function with $95 \%$ Confidence Limits


FILENAME: CYPC TRIAL V7.SAS - DATE: 23NOV17

## Cox Proportional Hazards Regression

## Cox Proportional Hazards

- K-M Curves are limited by not being able to control or adjust survival for other co-variates (only stratified analysis)
- Cox Regression is semi-parametric (you do not need to specify a probability distribution for survival times)
- Can easily incorporate time-dependent covariates
- Can use discrete and continuous time measures (you may only measure an outcome every year)


## Cox Proportional Hazards II

- Reminder
- Hazard Function quantifies the instantaneous risk that an event will occur at time $t$
- Key Assumption is proportional hazards
- survival curves for two strata (defined by any covariate you put in the model) must have hazard functions that are proportional over time (i.e. constant relative hazard)
- Test this by introducing an interaction with time for each covariate and testing if the interaction term is statistically significant


## PROC PHREG DATA = T7; <br> CLASS EARLY_AGE (REF="YOUNG"); <br> MODEL TIMESURVIVAL*CENSOS(1) = EARLY_AGE / RL; RUN;

/* note I have recoded early_age to be numeric */


## The PHREG Procedure

$$
\text { Type } 3 \text { Tests }
$$

| Effect | DF | Wald <br> Chi-Square | $\mathrm{Pr}>\mathrm{ChiSq}$ |
| :--- | :--- | ---: | ---: |
| EARLY_AGE | 2 | 163.2025 | $<.0001$ |

## Analysis of Maximum Likelihood Estimates

| Parameter | DF | Parameter Estimate | Standard Error |  | $\mathrm{Pr}>$ Chisq | Hazard Ratio | 95\% Hazard Confidence | Ratio <br> Limits | Label |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Chisquare |  |  |  |  |  |
| EARLY_AGE INFANT | 1 | 2.57624 | 0.20166 | 163.2015 | <. 0001 | 13.148 | 8.855 | 19.521 | EARLY_AGE INFANT |
| EARLY_AGE OLD | 1 | 0.74155 | 0.14893 | 24.7912 | <.0001 | 2.099 | 1.568 | 2.811 | EARLY_AGE OLD |

# PROC PHREG DATA = T7; <br> CLASS EARLY_AGE (REF="YOUNG") MALE (REF="0"); <br> MODEL TIMESURVIVAL*CENSOS(1) = EARLY_AGE MALE / RL; RUN; 

Type 3 Tests
Wald
Effect DF Chi-Square $\mathrm{Pr}>$ Chisq

| EARLY_AGE | 2 | 166.9010 | $<.0001$ |
| :--- | :--- | ---: | ---: |
| MALE | 1 | 6.9780 | 0.0083 |

Analysis of Maximum Likelihood Estimates

| Parameter | Parameter |  | Standard |  |  | Hazard Ratio | 95\% Hazard Ratio |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | OF |  |  | Chi-square | Pr $>$ Chisq |  |  |  | Label |
| EARLY_AGE INFANT | 1 | 2.61208 | 0.20220 | 166.8854 | <. 0001 | 13.627 | 9.169 | 20.255 | EARLY_AGE INFANT |
| EARLY_AGE OLD | 1 | 0.73262 | 0.14897 | 24.1858 | <. 0001 | 2.081 | 1.554 | 2.786 | EARLY_AGE OLD |
| MALE 1 | 1 | 0.37553 | 0.14216 | 6.9780 | 0.0083 | 1.456 | 1.102 | 1.924 | MALE 1 |

## Testing Proportionality Assumption

```
PROC PHREG DATA = T7;
CLASS EARLY_AGE (REF="YOUNG") MALE (REF="0");
MODEL TIMESURVIVAL*CENSOS(1) = EARLY_AGE MALE AGE_T MALE_T/RL;
AGE_T = EARLY_AGE*LOG(TIMESURVIVAL);
MALE_T = MALE*LOG(TIMESURVIVAL);
PROPORTIONALITY_TEST: TEST AGE_T, MALE_T;
RUN;
```

Type 3 Tests

|  | Wald <br> Effect |  |  |
| :--- | ---: | ---: | ---: |
| DF | Chi-Square |  |  | Pr > Chisq

Analysis of Maximum Likelihood Estimates

| Parameter | DF | Parameter Estimate | Standard Error | Chi-Square | Pr > Chisq | Hazard Ratio | 95\% Hazard Confidence | Ratio <br> Limits | Label |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| EARLY_AGE INFANT | 1 | 3.05714 | 0.40235 | 57.7340 | <. 0001 | 21.267 | 9.665 | 46.793 | EARLY AGE INFANT |
| EARLY_AGE OLD | 1 | 0.22946 | 0.42496 | 0.2915 | 0.5892 | 1.258 | 0.547 | 2.893 | EARLY_AGE OLD |
| MALE 1 | 1 | 0.18871 | 0.60235 | 0.0982 | 0.7541 | 1.208 | 0.371 | 3.933 | MALE 1 |
| AGE_T | 1 | 0.07987 | 0.06484 | 1.5175 | 0.2180 | 1.083 | 0.954 | 1.230 |  |
| MALE_T | 1 | 0.02923 | 0.09814 | 0.0887 | 0.7658 | 1.030 | 0.849 | 1.248 |  |

Linear Hypotheses Testing Results
Wald

| Label | Chi-Square | DF | $\mathrm{Pr}>$ Chisq |
| :--- | ---: | :---: | :---: |
| PROPORTIONALITY_TEST | 1.6654 | 2 | 0.4349 |

## Topics Covered

- Time-To-Event Data Analysis
- Introduction to time-to-event analysis
- Kaplan-Meier Curves
- Testing difference over strata
- Cumulative Incidence Curves
- Use of macro
- Introduction to Cox Proportional Hazards Modeling
- Testing proportional hazards assumption

